

Example 2. Assume that you want to estimate the mean $\mathbb{E}_{\mu^\dagger}[X]$ of some random variable X with respect to some unknown distribution μ^\dagger on the interval $[0, 1]$ based on the observation single data point $d_1 \approx 0.5$ up to resolution δ (i.e. we observe $d_1 \in B_\delta(x_1)$ with $x_1 = 0.5$). Consider the following two Bayesian models $\mu^a(\theta)$ and $\mu^b(\theta)$ on the unit interval $[0, 1]$, parameterized by $\theta \in (0, 1)$, with densities $f^a(\theta)$ and $f^b(\theta)$ given by

$$f^a(x, \theta) = (1 - \theta)\left(1 + \frac{1}{\theta}\right)(1 - x)^{\frac{1}{\theta}} + \theta\left(1 + \frac{1}{1-\theta}\right)x^{\frac{1}{1-\theta}},$$

$$f^b(x, \theta) = \begin{cases} f^a(x, \theta) \frac{1}{Z(\theta)} \left(\mathbb{1}_{\{x \notin (x_1 - \frac{\delta_c}{2}, x_1 + \frac{\delta_c}{2})\}} + 10^{-9} \mathbb{1}_{\{x \in (x_1 - \frac{\delta_c}{2}, x_1 + \frac{\delta_c}{2})\}} \right), & \text{if } \theta < 0.999, \\ f^a(x, \theta), & \text{if } \theta \geq 0.999, \end{cases}$$

where $Z(\theta)$ is a normalization constant close to one, chosen so that $\int_{[0,1]} f^b(x, \theta) dx = 1$. See Figure 1 for an illustration of these densities.

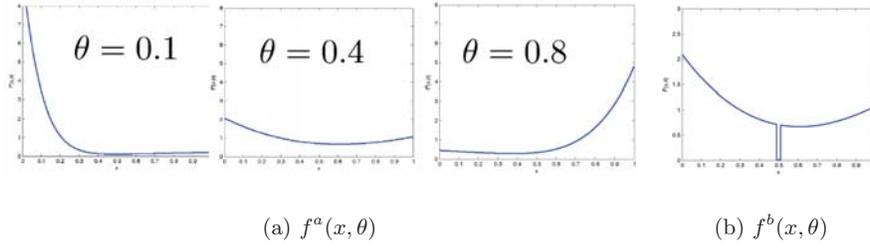


Figure 1: Illustration of the density $f^a(x, \theta)$ of model a and $f^b(x, \theta)$ of model b .

Observe that the density of model b is that of model a besides the small gap of width $\delta_c > 0$ created around the data point when $\theta < 0.999$ (see Figure 1); since the data point is fixed at $x_1 = \frac{1}{2}$, the total variation distance $d_{TV}(\mu^a(\theta), \mu^b(\theta))$ between the two models is, uniformly over $\theta \in (0, 1)$, less than a constant times δ_c . Assuming that the prior distribution on θ is the uniform distribution on $(0, 1)$, observe that the prior value of the quantity of interest $\mathbb{E}_\mu[X]$ under both models (a and b) is approximately $\frac{1}{2}$. Now, when θ is close to one (zero) then the density of model a puts most of its mass towards one (zero). Observe also that the density of model b behaves in a similar way, with the important exception that the probability of observing the data under model b is infinitesimally small for $\theta < 0.999$. Therefore, for $\delta < \delta_c$, the posterior value of the quantity of interest $\mathbb{E}_\mu[X]$ under model a is $\frac{1}{2}$ whereas it is close to one under model b (because, under the density $f^b(x, \theta)$, the probability of observing the data, up to resolution δ , is extremely unlikely if $\theta < 0.999$ when compared to the alternative $\theta \geq 0.999$). Observe also that a perturbed model c analogous to b can easily be constructed which leads to a posterior value close to zero. This simple example of brittleness under infinitesimal model perturbations is derived from the proof of theorem 6.4 of <http://arxiv.org/abs/1304.6772>, which shows that Bayesian posterior values

are generally brittle under infinitesimal perturbations of Bayesian models in TV and in Prokhorov metrics.